

# Hit and Miss Method

$$I = \int_a^b g(x) dx$$

Area of region  $S$  under  $g(x)$  curve.

$$f_{\hat{x}\hat{y}}(x, y) = \begin{cases} \frac{1}{c(b-a)} & \text{if } (x, y) \in \Omega \\ 0 & \text{if } (x, y) \notin \Omega \end{cases}$$

Probability  $p$  that  $(x, y)$  lies in  $S$  is:

$$p = \int_{\Omega} f_{\hat{x}\hat{y}}(x, y) dx dy = \frac{1}{c(b-a)} \int_S dx dy = \frac{I}{c(b-a)}$$

Assuming  $0 \leq g(x) \leq c$

Generate randomly  $N$  point  $(x_1, y_1), (x_2, y_2), \dots, (x_N, y_N)$

$\hat{N}_A$  number of points in  $S$

$\hat{N}_A$  follows the binomial distribution

$$\hat{\theta}_1 = \frac{c(b-a)}{N} \hat{N}_A$$

the mean value equals the integral:

$$\langle \hat{\theta}_1 \rangle = \frac{c(b-a)}{N} \langle \hat{N}_A \rangle = c(b-a)p = I$$

$\hat{\theta}_1$  is an *unbiased estimator* of  $I$ .

$$\sigma[\hat{\theta}_1] = \frac{c(b-a)}{N} \sqrt{Np(1-p)} = c(b-a) \sqrt{\frac{p(1-p)}{N}}$$

Replace  $p$  by its sample value:  $\hat{p} = \hat{N}_A/N$ .

$$I = \hat{\theta}_1 \pm \sigma[\hat{\theta}_1] = c(b-a)\hat{p} \pm c(b-a)\sqrt{\frac{\hat{p}(1-\hat{p})}{N}}$$

Relative error:

$$\frac{\sigma[\hat{\theta}_1]}{<\hat{\theta}_1>} = \sqrt{\frac{1-p}{pN}}$$

decreases for large  $p$ .

Take  $c = \max(g(x))$ .

```
subroutine mc1(g,a,b,c,n,r,s)
external g
na=0
do 1 i=1,n
u=ran_u()
v=ran_u()
if (g(a+(b-a)*u).gt.c*v) na=na+1
1 continue
p=real(na)/n
r=(b-a)*c*p
s=sqrt(p*(1.-p)/n)*c*(b-a)
return
end
```

## Sampling methods: Uniform Sampling

$$I = \int_a^b g(x) dx$$

$$I = \int_a^b (b-a)g(x) \frac{1}{b-a} dx \equiv (b-a) \int g(x) f_{\hat{x}}(x) dx$$

$$I = (b-a)E_{\mathbf{x}}[g(x)]$$

$\hat{\mathbf{x}}$  es  $\hat{U}(a, b)$ :

$$f_{\hat{x}}(x) = \begin{cases} \frac{1}{b-a} & \text{if } a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

Sample mean:

$$\hat{\mu}_N[g(x)] = \frac{1}{N} \sum_{i=1}^N g(x_i)$$

r.v.  $\hat{\theta}_2 = (b-a)\hat{\mu}_N$  from sampling:

$$\hat{\theta}_2 = (b-a) \frac{1}{N} \sum_{i=1}^N g(x_i)$$

$x_i, i = 1, 2, \dots, N$  uniformly distributed in the interval  $[a, b]$ .

$$\langle \theta_2 \rangle = I$$

$$I = \hat{\theta}_2 \pm \sigma[\hat{\theta}_2] = \hat{\theta}_2 \pm \frac{(b-a)\sigma[g(x)]}{\sqrt{N}}$$

$$\hat{\sigma}_N^2 = \left[ \frac{1}{N} \sum_{i=1}^N g(x_i)^2 - \left( \frac{1}{N} \sum_{i=1}^N g(x_i) \right)^2 \right]$$

```
subroutine mc2(g,a,b,n,r,s)
external g
r1=0.
s1=0.
do 1 i=1,n
u=ran_u()
g0=g(a+(b-a)*u)
r1=r1+g0
s1=s1+g0*g0
1 continue
r1=r1/n
s1=sqrt((s1/n-r1*r1)/n)
r=(b-a)*r1
s=(b-a)*s1
return
end
```

# Method of Importance Sampling

$$I = \int_a^b g(x) f_{\hat{\mathbf{x}}}(x) dx$$

$$I = E_{\mathbf{x}}[g(x)]$$

$\hat{\mathbf{x}}$  distributed according  $f_{\hat{\mathbf{x}}}(x)$ . Sample mean:

$$\hat{\mu}_N[g(x)] = \frac{1}{N} \sum_{i=1}^N g(x_i)$$

$x_i, i = 1, 2, \dots, N$  are values of the r.v.  $\hat{\mathbf{x}}$  distributed according to  $f_{\hat{\mathbf{x}}}(x)$

Unbiased estimator  $\langle \hat{\mu}_N[g(x)] \rangle = I$ .

$$I = \hat{\mu}_N[g(x)] \pm \sigma[\hat{\mu}_N[g(x)]]$$

$$I = \hat{\mu}_N[g(x)] \pm \frac{\sigma[g(x)]}{\sqrt{N}}$$

$$\sigma^2[g(x)] = \int f_{\hat{\mathbf{x}}} g(x)^2 dx - (\int f_{\hat{\mathbf{x}}} g(x) dx)^2$$

Use sample variance  $\hat{\sigma}_N^2$ :

$$\hat{\sigma}_N^2 = \left[ \frac{1}{N} \sum_{i=1}^N g(x_i)^2 - \left( \frac{1}{N} \sum_{i=1}^N g(x_i) \right)^2 \right]$$

How do we generate  $\hat{\mathbf{x}}_i$ ?

If  $u_i$  is a  $\hat{U}(0, 1)$  variable, then:

$$x_i = F_{\hat{\mathbf{x}}}^{-1}(u_i)$$

```
subroutine mc3(g,n,r,s)
external g
r=0.
s=0.
do 1 i=1,n
x=ran_f()
g0=g(x)
r=r+g0
s=s+g0*g0
1 continue
r=r/n
s=sqrt((s/n-r*r)/n)
return
end
```

## Efficiency of an integration method

Computer time needed to compute a given integral with a given error  $\epsilon$ .

So far:

$$\epsilon = \frac{\sigma}{\sqrt{N}}$$

$t$  is the computer time needed to add a contribution to the estimator

total computer time  $Nt \propto t\sigma^2$

Relative efficiency of method 1 and 2:

$$e_{12} = \frac{t_1\sigma_1^2}{t_2\sigma_2^2}$$

Uniform sampling is more efficient than hit and miss method.

Hit and miss:

$$\sigma_1^2 = c(b-a)I - I^2$$

Uniform sampling:

$$\sigma_2^2 = (b-a) \int g(x)^2 dx - I^2$$

$$\sigma_1^2 - \sigma_2^2 = (b-a) [cI - \int g(x)^2 dx]$$

The condition  $0 \leq g(x) \leq c$  implies:

$$\int g(x)^2 dx \leq c \int g(x) dx = cI$$

and  $\sigma_1 \geq \sigma_2$ . Furthermore, in general,  $t_1 > t_2$ , and it follows  $e_{12} > 1$ .

## Advantages and disadvantages of the Monte-Carlo integration

Very singular function

$n$ -dimensional integral, with high  $n$

$$\int_{\Omega} g(x_1, x_2, \dots, x_n) f_{\hat{\mathbf{x}}_1 \dots \hat{\mathbf{x}}_n}(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n \approx$$

$$\frac{1}{N} \sum_{i=1}^N g(x_1, x_2, \dots, x_n)$$

$(x_1, x_2, \dots, x_n)$  is a random vector distributed according the probability density function  $f_{\hat{\mathbf{x}}_1 \dots \hat{\mathbf{x}}_n}(x_1, x_2, \dots, x_n)$ .

Everything is valid for sums:

$$\sum_i g_i = \int g(x) f_{\hat{\mathbf{x}}}(x) dx$$

by considering a discrete random variable

$$f_{\hat{\mathbf{x}}}(x) = \sum_i p_i \delta(x - x_i)$$

and

$$g_i = g(x_i)$$